

Analysis of Variance – Multiple Regression

Model: $Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \varepsilon \quad \varepsilon \sim NID(0, \sigma^2) \quad p' = p + 1$

$$\mathbf{Y} = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} \quad \mathbf{X} = \begin{bmatrix} 1 & X_{11} & X_{12} & \cdots & X_{1p} \\ 1 & X_{21} & X_{22} & \cdots & X_{2p} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & X_{n1} & X_{n2} & \cdots & X_{np} \end{bmatrix} = [\mathbf{1} \quad \mathbf{X}^*] \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{bmatrix} = \begin{bmatrix} \beta_0 \\ \boldsymbol{\beta}^* \end{bmatrix} \quad \boldsymbol{\varepsilon} = \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon} = \beta_0 \mathbf{1} + \mathbf{X}^* \boldsymbol{\beta}^* + \boldsymbol{\varepsilon} \quad E\{\boldsymbol{\varepsilon}\} = \mathbf{0} \quad V\{\boldsymbol{\varepsilon}\} = \sigma^2 \mathbf{I} \quad E\{\mathbf{Y}\} = \mathbf{X}\boldsymbol{\beta} = \beta_0 \mathbf{1} + \mathbf{X}^* \boldsymbol{\beta}^* \quad V\{\mathbf{Y}\} = \sigma^2 \mathbf{I}$$

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \quad \hat{\mathbf{Y}} = \mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} = \mathbf{P}\mathbf{Y} \quad \mathbf{e} = \mathbf{Y} - \hat{\mathbf{Y}} = (\mathbf{I} - \mathbf{P})\mathbf{Y} \quad \bar{Y} = \bar{Y}\mathbf{1} = \left(\frac{1}{n}\right)\mathbf{J}\mathbf{Y}$$

$$V\{\hat{\boldsymbol{\beta}}\} = \sigma^2 (\mathbf{X}'\mathbf{X})^{-1} \quad V\{\hat{\mathbf{Y}}\} = \sigma^2 \mathbf{P} \quad V\{\mathbf{e}\} = \sigma^2 (\mathbf{I} - \mathbf{P})$$

Total (Corrected) Sum of Squares: $TSS = \sum_{i=1}^n (Y_i - \bar{Y})^2 = \mathbf{Y}' \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{Y}$

$$df_{\text{Tot}} = \text{rank} \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) = \text{tr} \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) = n - 1$$

Error (Residual) Sum of Squares: $SSE = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 = \mathbf{Y}'(\mathbf{I} - \mathbf{P})\mathbf{Y}$

$$df_{\text{Err}} = \text{tr}(\mathbf{I} - \mathbf{P}) = \text{tr}(\mathbf{I} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}') = \text{tr}(\mathbf{I}) - \text{tr}(\mathbf{X}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}) = n - p'$$

Regression Sum of Squares: $SSR = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 = \mathbf{Y}' \left(\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{Y} \quad df_{\text{Reg}} = \text{tr} \left(\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J} \right) = p' - 1 = p$

Distribution of $\frac{SSE}{\sigma^2}$: $\frac{SSE}{\sigma^2} = \mathbf{Y}' \left(\frac{\mathbf{I} - \mathbf{P}}{\sigma^2} \right) \mathbf{Y} = \mathbf{Y}' \mathbf{A}_E \mathbf{Y} \Rightarrow \mathbf{A}_E V\{\mathbf{Y}\} \mathbf{A}_E V\{\mathbf{Y}\} = \left(\frac{\mathbf{I} - \mathbf{P}}{\sigma^2} \right) \sigma^2 \mathbf{I} \left(\frac{\mathbf{I} - \mathbf{P}}{\sigma^2} \right) \sigma^2 \mathbf{I} = \left(\frac{\mathbf{I} - \mathbf{P}}{\sigma^2} \right) \sigma^2 \mathbf{I}$

$$\Rightarrow \frac{SSE}{\sigma^2} \sim \chi^2 \quad \text{with } df_{\text{Err}} = \text{rank} \left(\frac{\mathbf{I} - \mathbf{P}}{\sigma^2} \right) = n - p' \quad \text{and } \Omega_{\text{Err}} = \frac{1}{2\sigma^2} \boldsymbol{\beta}' \mathbf{X}' (\mathbf{I} - \mathbf{P}) \mathbf{X} \boldsymbol{\beta} = \mathbf{0}$$

Distribution of $\frac{SSR}{\sigma^2}$: $\frac{SSR}{\sigma^2} = \mathbf{Y}' \left(\frac{\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J}}{\sigma^2} \right) \mathbf{Y} = \mathbf{Y}' \mathbf{A}_R \mathbf{Y}$

$$\Rightarrow \mathbf{A}_R V\{Y\} \mathbf{A}_R V\{Y\} = \left(\frac{\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J}}{\sigma^2} \right) \sigma^2 \mathbf{I} \left(\frac{\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J}}{\sigma^2} \right) \sigma^2 \mathbf{I} = \left(\frac{\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J}}{\sigma^2} \right) \sigma^2 \mathbf{I}$$

$$\Rightarrow \frac{SSR}{\sigma^2} \sim \chi^2 \text{ with } df_{\text{Reg}} = \text{rank} \left(\frac{\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J}}{\sigma^2} \right) = p \text{ and}$$

$$\Omega_{\text{Reg}} = \frac{1}{2\sigma^2} \boldsymbol{\beta}' \mathbf{X}' \left(\mathbf{P} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{X} \boldsymbol{\beta} = \frac{1}{2\sigma^2} \boldsymbol{\beta}' \mathbf{X}' \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{X} \boldsymbol{\beta}$$

Now write: $\frac{1}{2\sigma^2} \boldsymbol{\beta}' \mathbf{X}' \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{X} \boldsymbol{\beta} = \frac{1}{2\sigma^2} [\beta_0 \quad \boldsymbol{\beta}^*]' \begin{bmatrix} \mathbf{1}' \\ \mathbf{X}^* \end{bmatrix} \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \begin{bmatrix} \mathbf{1} & \mathbf{X}^* \end{bmatrix} \begin{bmatrix} \beta_0 \\ \boldsymbol{\beta}^* \end{bmatrix} =$

$$\frac{1}{2\sigma^2} (\beta_0 \mathbf{1}' + \boldsymbol{\beta}^{*'} \mathbf{X}^{*'}) \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) (\beta_0 \mathbf{1} + \mathbf{X}^* \boldsymbol{\beta}^*)$$

Note that: $\mathbf{1}' \mathbf{I} = \mathbf{1}'$, $\mathbf{1}' \left(\frac{1}{n}\right)\mathbf{J} = \mathbf{1}'$ and $\mathbf{I} \mathbf{1} = \mathbf{1}$, $\left(\frac{1}{n}\right)\mathbf{J} \mathbf{1} = \mathbf{1} \Rightarrow \mathbf{1}' \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) = \mathbf{0}$ and $\left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{1} = \mathbf{0}$

$$\Rightarrow \Omega_{\text{Reg}} = \frac{1}{2\sigma^2} \boldsymbol{\beta}^{*'} \mathbf{X}^{*'} \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{X}^* \boldsymbol{\beta}^*$$

Now, consider for vector \mathbf{w} :

$$\mathbf{w}' \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{w} = \mathbf{w}' \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right)' \left(\mathbf{I} - \left(\frac{1}{n}\right)\mathbf{J} \right) \mathbf{w} = (\mathbf{w} - \bar{\mathbf{w}})' (\mathbf{w} - \bar{\mathbf{w}})$$

$$= \begin{bmatrix} w_1 - \bar{w} & w_2 - \bar{w} & \cdots & w_n - \bar{w} \end{bmatrix} \begin{bmatrix} w_1 - \bar{w} \\ w_2 - \bar{w} \\ \vdots \\ w_n - \bar{w} \end{bmatrix} = \sum_{i=1}^n (w_i - \bar{w})^2 = 0 \Leftrightarrow \mathbf{w} = \bar{\mathbf{w}}$$

$$\text{In the case of SSR: } \mathbf{X}^* \boldsymbol{\beta}^* = \begin{bmatrix} \beta_1 X_{11} + \dots + \beta_p X_{1p} \\ \beta_1 X_{21} + \dots + \beta_p X_{2p} \\ \vdots \\ \beta_1 X_{n1} + \dots + \beta_p X_{np} \end{bmatrix} \quad \overline{\mathbf{X}^* \boldsymbol{\beta}^*} = \begin{bmatrix} \beta_1 \bar{X}_1 + \dots + \beta_p \bar{X}_p \\ \beta_1 \bar{X}_1 + \dots + \beta_p \bar{X}_p \\ \vdots \\ \beta_1 \bar{X}_1 + \dots + \beta_p \bar{X}_p \end{bmatrix} \quad \mathbf{X}^* \boldsymbol{\beta}^* - \overline{\mathbf{X}^* \boldsymbol{\beta}^*} = \begin{bmatrix} \sum_{j=1}^p \beta_j (X_{1j} - \bar{X}_j) \\ \sum_{j=1}^p \beta_j (X_{2j} - \bar{X}_j) \\ \vdots \\ \sum_{j=1}^p \beta_j (X_{nj} - \bar{X}_j) \end{bmatrix}$$

$$\Rightarrow \Omega_{\text{Reg}} = 0 \Leftrightarrow \sum_{j=1}^p \beta_j (X_{ij} - \bar{X}_j) = 0 \quad i = 1, \dots, n$$

$$\text{Case 1: } \beta_1 = \dots = \beta_p = 0 \Rightarrow \sum_{j=1}^p \beta_j (X_{ij} - \bar{X}_j) = 0 \quad i = 1, \dots, n$$

$$\text{Case 2: } \beta_k \neq 0, \beta_{k'} = 0 \quad k' \neq k \Rightarrow \beta_k (X_{ik} - \bar{X}_k) = 0 \quad i = 1, \dots, n \quad \text{which means no variation in } X_k$$

$$\text{Case 3: } \beta_{k_1}, \dots, \beta_{k_r} \neq 0 \Rightarrow \sum_{j=1}^r \beta_{k_j} (X_{ik_j} - \bar{X}_{k_j}) = 0 \quad i = 1, \dots, n$$

$$\Rightarrow X_{ik_j} = - \sum_{\substack{j'=1 \\ j' \neq j}}^r \frac{\beta_{k_{j'}}}{\beta_{k_j}} (X_{ik_{j'}} - \bar{X}_{k_{j'}}) + \bar{X}_{k_j} = - \sum_{\substack{j'=1 \\ j' \neq j}}^r \frac{\beta_{k_{j'}}}{\beta_{k_j}} X_{ik_{j'}} + \sum_{j'=1}^r \frac{\beta_{k_{j'}}}{\beta_{k_j}} \bar{X}_{k_{j'}} \quad i = 1, \dots, n$$

which implies linear dependencies among the columns of $\mathbf{X} = [\mathbf{1} \quad \mathbf{X}^*]$.